JAMA | Original Investigation

Association of Stroke Among Adults Aged 18 to 49 Years With Long-term Mortality

Merel Sanne Ekker, MD; Jamie Inge Verhoeven, Bsc; Ilonca Vaartjes, PhD; Wilhelmus Martinus Tim Jolink, MD; Catharina Johanna Maria Klijn, MD, PhD; Frank-Erik de Leeuw, MD, PhD

IMPORTANCE Stroke remains the second leading cause of death worldwide. Approximately 10% to 15% of all strokes occur in young adults. Information on prognosis and mortality specifically in young adults is limited.

OBJECTIVE To determine short- and long-term mortality risk after stroke in young adults, according to age, sex, and stroke subtype; time trends in mortality; and causes of death.

DESIGN, SETTING, AND PARTICIPANTS Registry- and population-based study in the Netherlands of 15 527 patients aged 18 to 49 years with first stroke between 1998 and 2010, and follow-up until January 1, 2017. Patients and outcomes were identified through linkage of the national Hospital Discharge Registry, national Cause of Death Registry, and the Dutch Population Register.

EXPOSURES First stroke occurring at age 18 to 49 years, documented using *International Classification of Diseases, Ninth Revision, and International Statistical Classification of Diseases and Related Health Problems, Tenth Revision,* codes for ischemic stroke, intracerebral hemorrhage, and stroke not otherwise specified.

MAIN OUTCOMES AND MEASURES Primary outcome was all-cause cumulative mortality in 30-day survivors at end of follow-up, stratified by age, sex, and stroke subtype, and compared with all-cause cumulative mortality in the general population.

RESULTS The study population included 15 527 patients with stroke (median age, 44 years [interquartile range, 38-47 years]; 53.3% women). At end of follow-up, a total of 3540 cumulative deaths had occurred, including 1776 deaths within 30 days after stroke and 1764 deaths (23.2%) during a median duration of follow-up of 9.3 years (interquartile range, 5.9-13.1 years). The 15-year mortality in 30-day survivors was 17.0% (95% CI, 16.2%-17.9%). The standardized mortality rate compared with the general population was 5.1 (95% CI, 4.7-5.4) for ischemic stroke (observed mortality rate 12.0/1000 person-years [95% CI, 11.2-12.9/1000 person-years]; expected rate, 2.4/1000 person-years; excess rate, 9.6/1000 person-years) and the standardized mortality rate for intracerebral hemorrhage was 8.4 (95% CI, 7.4-9.3; observed rate, 18.7/1000 person-years [95% CI, 16.7-21.0/1000 person-years]; expected rate, 2.2/1000 person-years; excess rate, 16.4/1000 person-years].

CONCLUSIONS AND RELEVANCE Among young adults aged 18 to 49 years in the Netherlands who were 30-day survivors of first stroke, mortality risk compared with the general population remained elevated up to 15 years later.

JAMA. 2019;321(21):2113-2123. doi:10.1001/jama.2019.6560 Published online May 23, 2019. + Supplemental content

Author Affiliations: Radboud University Medical Centre, Donders Institute for Brain, Cognition, and Behaviour, Department of Neurology, Nijmegen, the Netherlands (Ekker, Verhoeven, Klijn, de Leeuw); Julius Center for Health Sciences and Primary Care, University Medical Center Utrecht, Utrecht, the Netherlands (Vaartjes); Department of Neurology and Neurosurgery, Brain Center Rudolf Magnus, University Medical Center Utrecht, Utrecht, the Netherlands (Jolink, Klijn).

Corresponding Author: Frank-Erik de Leeuw, MD, PhD, Department of Neurology (935), Radboud University Medical Centre, PO Box 9101, 6500 HB Nijmegen, the Netherlands (FrankErik.deLeeuw@radboudumc.nl). S troke is the second leading cause of death worldwide. In 2013, more than 10 million people experienced a stroke, and more than 6 million died.¹⁻³ Approximately 10% to 15% of all strokes occur in young adults aged 18 to 49 years.^{4,5} Information on the risk of death in this subgroup is limited.⁶

Previous studies of mortality after stroke in young adults were often small, hospital based, had limited periods of followup, or included patients who had their stroke between 1980 and 2000.⁷⁻⁹ Over the past decades, both acute treatment and secondary prevention have improved. Because of the small number of patients younger than age 50 years with stroke included in previous studies, it has not been possible to estimate mortality according to age, sex, and stroke subtypes.^{7,10-15}

This study aimed to investigate case fatality and cumulative 1-year, 5-year, 10-year, and 15-year mortality and trends over time of first stroke in young adults aged 18 to 49 years, stratified by age, sex, and stroke subtype; excess mortality after stroke compared with the general population; and causes of death.

Methods

The Medical Ethics Review Committee Arnhem/Nijmegen assessed the protocol and waived the requirement for ethical review and for patient consent due to the use of deidentified data. All analyses were performed in a secured environment of Statistics Netherlands according to Dutch privacy legislation.

Exposure

We constructed a nationwide cohort of patients aged 18 to 49 years with first stroke (ischemic stroke, intracerebral hemorrhage, or stroke not otherwise specified) through linkage of the Dutch nationwide hospital registry (Hospital Discharge Registry [HDR]) and Dutch population registry from January 1, 1998, to January 1, 2011, and the National Cause of Death Registry (CDR) from January 1, 1998, to January 1, 2017, by using International Classification of Diseases, Ninth Revision (ICD-9) and International Statistical Classification of Diseases and Related Health Problems, Tenth Revision (ICD-10) codes for stroke (ischemic stroke, intracerebral hemorrhage, and stroke not otherwise specified; World Health Organization International Classification of Diseases) (eTable 1 in the Supplement). We did not include transient ischemic attacks (TIAs) and subarachnoid hemorrhage (SAH). Details of these registries and linkage procedures have been previously described.14,16,17

ICD-9 and *ICD-10* codes for the identification of stroke have been proven to be reliable in studies of patients with stroke of all ages.^{18,19} For this study, we specifically assessed the accuracy of *ICD-10* codes for young adults (aged 18-49 years). We checked final diagnoses in medical records against the attached *ICD-10* code at discharge of 569 patients admitted in 2 university medical centers and 1 large general hospital between 1995 and 2017. For ischemic stroke, the *ICD* code was correct in 90.4% of cases (n = 301), for intracerebral hemorrhage in 86.3% of cases (n = 183),

Key Points

Question In young adults aged 18 to 49 years, what is the age- and sex-specific case fatality and long-term mortality associated with stroke?

Findings In this Dutch register-based cohort study that included 15 527 patients who in the years 1998-2010 had a first stroke at age 18 to 49 years, cumulative 15-year mortality among 30-day survivors was 13.3 per 1000 person-years compared with an expected mortality of 2.4 per 1000 person-years in the general population, an excess mortality of 10.9 per 1000 person-years.

Meaning Mortality risk 15 years after stroke among young adults aged 18 to 49 years who were 30-day survivors remained elevated.

and for stroke not otherwise specified in 87.1% of cases (n = 85) (eTable 2 in the Supplement).

For all individuals, the Charlson Comorbidity Index (CCI) score at the time of the index event was estimated based on previously documented hospital admissions. The CCI is a score from 1 to 6 based on 19 primary medical conditions, such as congestive heart failure, diabetes, malignancies, and other organ dysfunction, and has previously been shown to be a valid tool to predict outcome.²⁰

Outcomes

The primary outcome was all-cause cumulative mortality at the end of follow-up, stratified for age, sex, and stroke subtype in 30-day survivors. The secondary outcomes were case fatality and cumulative 1-year, 5-year, 10-year, and 15-year mortality in the 30-day survivors. Other secondary outcomes were the annual risk of death, excess mortality after stroke compared with the general population, time trends of case fatality, 1-year and 5-year mortality in 30-day survivors after first stroke in young adults, and causes of death, all stratified by age, sex, and stroke subtype.

Out-of-hospital deaths were identified if an ICD-10 code for stroke was registered in the CDR without previous hospitalizations in the HDR for this individual. Individuals who emigrated were censored. The number of hospitals participating in the HDR declined from 2005 to 2010, leading to an increasing number of missing records, varying from 1.1% to 14%, resulting in some missing index strokes. Date and cause of death were retrieved from the CDR. In the Netherlands, all deaths are recorded in the CDR by Statistics Netherlands. Cause of death was missing for 1 patient (0.05%); this patient was excluded from the analysis. Data regarding other variables, age, sex, CCI, and length of stay were complete for all patients. Follow-up after first stroke between 1998 and 2010 was defined as the time until death or the end of follow-up (January 1, 2017), whichever occurred first. Case fatality was defined as death occurring within 30 days after stroke. For survival analysis, only survivors beyond these 30 days were included. Causes of death were analyzed by ICD-10 codes.

Statistical Analysis

We used Kaplan-Meier analysis to estimate the risk of death for all-cause stroke, ischemic stroke, intracerebral hemorrhage, and stroke not otherwise specified for men and women separately. We calculated the number of person-years at risk for each individual patient from date of stroke to the date of death or until January 1, 2017, whichever occurred first. We censored the Kaplan-Meier curves at 16 years after the index event because thereafter, the number of person-years at risk became too small to provide reliable estimates of mortality. We calculated 1-year, 5-year, 10-year, and 15-year cumulative mortality by sex and stroke subtype. Survival curves were compared between men and women using log-rank analysis.

For comparison of mortality in young adults with stroke vs mortality in the general population, we used mortality data of the general Dutch population matched by sex, age, and calendar year.¹⁷ The annual risk of the observed mortality after stroke, as well as the expected mortality in the general population, was calculated using the formula: $1 - ([1 - Ic]^{[1/n]})$, where n is the number of years after the index event and Ic is the cumulative mortality at n years, obtained by Kaplan-Meier analysis.

We calculated standardized mortality ratios (SMRs) by dividing the observed deaths in the cohort by the expected deaths of their peers from the general population for each stroke subtype, for both sexes and for different age categories (18-29, 30-39, and 40-49 years). The expected matched mortality rates were retrieved from the worldwide Human Mortality Database (http://www.mortality.org).²¹ The 95% CIs were calculated by assuming a Poisson distribution. Additionally, for each subgroup, we calculated an absolute excess number of deaths by taking the difference between the observed and expected deaths, divided by the person-years at risk. Differences of the SMRs by sex and stroke subtype were evaluated by tests of interactions through bivariable and multivariable linear regression modeling.

We assessed the association of age at index event, sex, CCI score, and length of stay with the risk of long-term mortality, stratified for stroke subtypes, through Cox proportional hazard models, and expressed the associations as hazard ratios (HRs) with 95% CIs. After bivariable analysis, all 4 variables were simultaneously entered to obtain a multivariable model. Assumption of proportionality in the Cox regression model was evaluated graphically assessing the log(-log[Survival]) plots for all covariates. We found no indication of violating the assumption. In addition, we plotted the scaled Schoenfeld residuals against time, which confirmed proportionality. We calculated time trends in case fatality and 1-year and 5-year mortality rates in yearly average percentage change (APC) and tested change over time with linear regression analysis. We used R^2 to assess the goodness of fit of the linear regression analyses. Time trends regarding 10-year and 15-year mortality were not calculated because we did not have a complete 10-year and 15-year follow-up period for all inclusion years. Also, time trends were calculated for mean length of stay and tested through linear regression.

In the Netherlands, cause of death is established among all patients who die (both inside or outside the hospital) by a coroner or physician who is required by law to complete the death certificate. This death certificate is then sent to coders employed by Statistics Netherlands who assign an *ICD-10* code for the primary (underlying disease) and secondary (possible complications, such as pneumonia) causes of death accordingly. Studies on the reliability of these *ICD-10* codes found an intercoder agreement of 78% and intracoder agreement of 89%. The highest reliability was found for major causes of death (malignancy and acute myocardial infarction).²² We categorized causes of death based on the *ICD-10* codes for primary cause in 30-day survivors (eTable 3 in the Supplement). The proportion of causes of death were compared through χ^2 analyses. We tested for interactions between stroke subtype and different causes of death separately using a Poisson regression model, with adjustment for sex through adding this as a covariate to a multivariable Poisson model.

Two-sided *P* values of .05 were considered statistically significant. For the SMR analyses, we set the threshold for significance to a *P* value of .005 after Bonferroni adjustment for 10 subgroup analyses.

Data were analyzed with SPSS Software version 22 (IBM), R version 3.22 (packages rateratio.test, survival, survminer; R Project for Statistical Computing), Stata version 12 (StataCorp), and Microsoft Office Excel 2007.

Results

We identified 15 257 young adults with a stroke (53.3% women; median age, 44 years [interquartile range, 38-47 years]). A total of 8444 (55.3%) had ischemic stroke, 3077 (20.2%) had intracerebral hemorrhage, and 3736 (24.5%) had a stroke not otherwise specified. Less than 1% of strokes were out-ofhospital deaths, and most had no comorbidity (n = 12 803, 83.9%) (**Table 1**). The median duration of follow-up was 9.3 years (interquartile range, 5.9-13.1 years; Table 1).

Following any stroke, 1776 patients (11.6%) died in the first 30 days. Case fatality after ischemic stroke was 7.4% (n = 629) (7.6% [n = 291] in men, 7.4% [n = 338] in women); after intracerebral hemorrhage, 32.3% (29.8% [n = 991] in men, 34.8% [n = 519] in women); and after stroke not otherwise specified, 4.2% (3.8% [n = 65] in men, 4.4% [n = 91] in women). This resulted in a cohort of 13 481 30-day survivors of any stroke (7815 ischemic stroke, 2086 intracerebral hemorrhage, and 3580 stroke not otherwise specified; eTable 4 in the Supplement).

Cumulative Mortality

At end of follow-up, a total of 3540 patients (23.2%) had died. In 30-day survivors, the cumulative mortality after any stroke increased from 3.1% (95% CI, 2.8%-3.4%) after 1 year to 7.0% (95% CI, 6.6%-7.4%) at 5 years, 11.5% (95% CI, 11.0%-12.1%) at 10 years, and 17.0% (95% CI, 16.2%-17.9%) after 15 years (**Figure**, A, C, E, and G; eTable 5 in the Supplement).

Sex-Specific Annual Risk of Death by Stroke Subtype

The annual risk of death of 30-day survivors was highest in the first years after stroke and then stabilized (Figure, B, D, F, and H). After ischemic stroke, the risk of death after 1 year was 2.0% (95% CI, 1.8%-2.3%) for men and 1.6% (95% CI, 1.4%-1.8%) for women, and after 10 years decreased to an

Table 1. Demographics Including Age, Sex, Comorbidity, and Follow-up of Patients With Stroke Aged 18-49 Years

	Stroke, No. (%)											
	Any			Ischemic			Intracerebral Hemorrhage			Not Otherwise Specified		
	Total	Men	Women	Total	Men	Women	Total	Men	Women	Total	Men	Women
Patients	15 257 (100)	7127 (46.7)	8130 (53.3)	8444 (100)	3851 (45.6)	4593 (54.4)	3077 (100)	1585 (51.5)	1492 (48.5)	3736 (100)	1691 (45.3)	2045 (54.7)
Age, mean (SD), y	41.8 (6.8)	42.3 (6.5)	41.4 (7.0)	42.0 (6.6)	42.6 (6.1)	41.4 (7.0)	40.7 (7.5)	40.7 (7.5)	40.8 (7.5)	42.2 (6.4)	42.9 (6.0)	41.6 (6.7)
Charlson Comorbidity Index score ^a												
0	12 803 (83.9)	5951 (83.5)	6852 (84.3)	7178 (85.0)	3262 (84.7)	3916 (85.2)	2539 (82.5)	1307 (82.4)	1232 (82.6)	3086 (82.6)	1382 (81.7)	1704 (83.3)
1	1694 (11.1)	810 (11.4)	884 (10.8)	879 (10.4)	397 (10.3)	482 (10.5)	353 (11.5)	189 (11.9)	164 (10.9)	462 (12.4)	223 (13.2)	238 (11.6)
2	574 (3.8)	275 (3.9)	299 (3.7)	299 (3.5)	152 (3.9)	147 (3.2)	137 (4.5)	63 (4.0)	74 (5.0)	138 (3.7)	69 (4.1)	78 (3.8)
≥3	186 (1.2)	91 (1.3)	95 (1.2)	88 (1.0)	40 (1.0)	48 (1.0)	48 (1.6)	26 (1.6)	22 (1.4)	50 (1.3)	25 (1.5)	25 (1.2)
Duration of follow-up ^b												
Follow-up, median (IQR), y	9.3 (5.9-13.1)	9.2 (5.8-12.9)	9.5 (6.0-13.3)	9.5 (6.3-13.2)	9.3 (6.2-12.9)	9.7 (6.4-13.4)	6.9 (0.0-11.9)	7.1 (0.0-11.9)	6.7 (0.0-11.8)	10.3 (7.0-13.5)	10.2 (6.9-13.3)	13.7 (7.2-13.7)
>5 y	12 541 (82.2)	5789 (81.2)	6752 (83.0)	7355 (87.1)	3314 (86.1)	4041 (88.0)	1842 (59.9)	969 (61.1)	873 (58.5)	3344 (89.5)	1509 (89.2)	1838 (89.9)
>10 y	6928 (45.4)	3138 (44.0)	3790 (46.6)	3903 (46.2)	1716 (44.6)	2187 (47.6)	1059 (34.4)	549 (34.6)	510 (34.2)	1966 (52.6)	873 (51.6)	1093 (53.4)

Abbreviation: IQR, interquartile range

^a Charlson Comorbidity Index score ranges from 0 to 6, with higher scores indicating more comorbidity.

^b Duration of follow-up is defined as the time, in years, between event and death or end of study, whichever occurred first.

annual risk of 1.4% (95% CI, 1.2%-1.6%) for men and 0.9% (95% CI, 0.7%-1.1%) for women.

For intracerebral hemorrhage, the annual risk of death was 4.8% (95% CI, 4.4%-5.1%) for men and 3.9% (95% CI, 3.6%-4.3%) for women after 1 year and after 10 years decreased to an annual risk of 1.9% (95% CI, 1.7%-2.1%) for men and 1.4% (95% CI, 1.2%-1.6%) for women (Figure, B, D, F, and H).

Long-term Mortality in 30-Day Survivors of Stroke Compared With the General Population

The SMR for young adults with any stroke compared with peers from the general population matched by sex, age, and calendar year was 5.6 (95% CI, 5.3-5.9). The observed mortality rate was 13.3 per 1000 person-years (95% CI, 12.6-14.0) vs an expected rate of 2.4 per 1000 person-years, with an excess rate of 10.9 deaths per 1000 person-years. The SMR for ischemic stroke was 5.1 (95% CI, 4.7-5.4), with an observed mortality rate of 12.0 per 1000 person-years (95% CI, 11.2-12.9) and expected mortality rate of 2.4 per 1000 person-years and an excess rate of 9.6 per 1000 person-years. For intracerebral hemorrhage, the SMR was 8.4 (95% CI, 7.4-9.3; observed mortality rate, 18.7/1000 person-years [95% CI, 16.7-21.0]; expected rate, 2.2/1000 person-years; excess rate, 16.5/1000 person-years). The SMR for stroke not otherwise specified was 5.2 (95% CI, 4.7-5.8; observed rate, 12.9 [95% CI, 11.7-14.3]; expected rate, 2.5/1000 person-years; excess rate, 10.5/1000 person-years). There were no significant differences in SMR between men and women (bivariable linear regression model, *P* = .74; multivariable model including stroke subtype and age groups, P = .07). Results stratified for age groups and sex are summarized in Table 2.

Association of Age, Sex, Comorbidity, and Length of Stay With the Risk of Death After Index Stroke

In bivariable Cox regression analysis for any stroke, age was associated with mortality during follow-up of 30-day survivors (35-39 years: hazard ratio [HR], 2.4 [95% CI, 1.5-3.8], P for interaction < .001; 40-44 years: HR, 2.7 [95% CI, 1.7-4.2], P for interaction < .001; 45-49 years: HR, 3.8 [95% CI, 2.4-5.9], *P* for interaction < .001; all compared with the reference group aged 18-24 years). Age groups 25-29 and 30-34 were not significantly associated with a higher risk of mortality. Male sex was associated with a higher risk of mortality, with an HR of 1.5 (95% CI, 1.4-1.7; P for interaction < .001), as well as a CCI score above 0 (CCI of 1: HR 2.1 [95% CI, 1.9-2.4], P for interaction < .001; CCI of 2: HR, 6.5 [95% CI, 5.6-7.5], *P* for interaction < .001; CCI of 3 or higher: HR, 9.4 [95% CI, 7.5-11.7], *P* for interaction < .001). In addition, length of hospital stay longer than 14 days was associated with mortality (HR, 2.2 [95% CI, 1.9-2.6]; P for interaction < .001). In the multivariable Cox regression model, all of these associations remained significant. Table 3 shows HRs of the bivariable and multivariable Cox regressions models specified per stroke subtype.

Time Trends in Case Fatality, 1-Year and 5-Year Mortality, and Length of Stay for Any Stroke and for Stroke Subtypes

Case fatality after any stroke decreased from 15.2% (n = 160) in 1998 to 6.9% (n = 86) in 2010 (yearly APC, -5.5% [95% CI, -13.1% to 2.0%]; P < .001; $R^2 = 0.9$), from 8.4% (n = 47) in 1998 to 4.9% (n = 38) in 2010 for ischemic stroke (yearly APC, -1.3% [95% CI, -17.1% to 14.6%]; P < .001; $R^2 = 0.6$), from 37.9% (n = 89) in 1998 to 21.1% (n = 44) in 2010 for intracerebral

Figure. Cumulative Mortality and Annual Mortality Over Time in 30-Day Stroke Survivors



years (7.2-13.7) for ischemic stroke (C); 9.9 years (6.6-13.6) and 10.4 years (6.9-14.3)

between men and women. Shaded regions indicate 95% Cls.

	Total	Patient-Years	Observed	Observed Deaths per 1000 Person-Years (95% CI)	Expected	Expected Deaths per 1000 Person-Vears	Excess Rate per 1000 Person-Vears ^b	Standardized Mortality Rate
Any Stroke	Totat		Deatilis	(55% CI)	Deatilis	reison-rears	reison-reals	(55% CI)
Total	13 481	102 184	1356	13.3 (12.6-14.0)	242.8	2.4	10.9	5.6 (5.3-5.9)
Men	6299	46 937	756	16.1 (15.0-17.3)	131.1	2.8	13.3	5.8 (5.4-6.2)
Women	7182	55 247	600	10.9 (10.0-11.8)	111.8	2.0	8.8	5.4 (4.9-5.8)
18-29 v	928	7477	52	7.0 (5.3-9.1)	3.3	0.4	6.5	15.8 (11.6-20.4)
Men	340	2635	28	10.6 (7.4-15.4)	1.6	0.6	10.0	18.1 (11.7-25.2)
Women	588	4842	24	5.0 (3.3-7.4)	1.7	0.4	4.6	13.8 (8.6-19.5)
30-39 v	2951	23 423	228	9.7 (8.6-11.1)	25.2	1.1	8.7	9.1 (7.9-10.3)
Men	1285	9958	128	12.9 (10.8-15.3)	12.3	1.2	11.6	10.4 (8.6-12.3)
Women	1666	13 465	100	7.4 (6.1-9.0)	12.9	1.0	6.5	7.8 (6.3-9.3)
40-49 v	9602	71 284	1076	15.1 (14.2-16.0)	214.4	3.0	12.1	5.0 (4.7-5.3)
Men	4674	34 344	600	17.5 (16.1-18.9)	117.2	3.4	14.1	5.1 (4.7-5.5)
Women	4928	36 939	476	12.9 (11.8-14.1)	97.1	2.6	10.3	4.9 (4.5-5.3)
Ischemic Stro	oke		-		-			
Total	7815	58 668	704	12.0 (11.2-12.9)	139.1	2.4	9.6	5.1 (4.7-5.4)
Men	3560	26 228	382	14.6 (13.2-16.1)	73.5	2.8	11.8	5.2 (4.7-5.7)
Women	4255	32 440	322	9.9 (8.9-11.1)	65.6	2.0	7.9	4.9 (4.4-5.5)
18-29 v	487	3903	26	6.7 (4.5-9.8)	1.7	0.4	6.2	15.6 (10.2-21.6)
Men	162	1247	10	8.0 (4.3-14.9)	0.7	0.6	7.4	13.7 (5.5-23.2)
Women	325	2656	16	6.0 (3.7-9.8)	0.9	0.4	5.7	17.2 (9.7-25.7)
30-39 v	1667	13 414	100	7.5 (6.1-9.1)	14.5	1.1	6.4	6.9 (5.6-8.3)
Men	696	5510	52	9.4 (7.2-12.4)	6.8	1.2	8.2	7.6 (5.6-9.8)
Women	971	7904	48	6.1 (4.6-8.1)	7.7	1.0	5.1	6.3 (4.6-8.1)
40-49 v	5661	41 351	578	14.0 (12.9-15.2)	123.0	3.0	11.0	4.7 (4.3-5.1)
Men	2702	19471	320	16.4 (14.7-18.3)	66.0	3.4	13.1	4.9 (4.3-5.4)
Women	2959	21 880	258	11.8 (10.4-13.3)	57.0	2.6	9.2	4.5 (4.0-5.1)
Intracerebral	Hemorrhage			. ,				. ,
Total	2086	15 560	291	18.7 (16.7-21.0)	34.8	2.2	16.5	8.4 (7.4-9.3)
Men	1113	8158	174	21.3 (18.4-24.7)	20.5	2.5	18.8	8.5 (7.3-9.8)
Women	973	7402	117	15.8 (13.2-18.9)	14.4	1.9	13.9	8.2 (6.7-9.7)
18-29 y	228	1835	17	9.3 (5.8-14.9)	0.8	0.5	8.8	20.4 (10.8-31.1)
Men	109	866	11	12.7 (7.1-22.9)	0.5	0.6	12.1	22.0 (10.0-35.9)
Women	119	969	6	6.2 (2.8-13.8)	0.3	0.4	5.9	17.9 (6.0-32.9)
30-39 y	515	3864	72	18.6 (14.8-23.4)	4.1	1.1	17.6	17.7 (13.8-21.9)
Men	279	2042	46	22.5 (16.9-30.0)	2.4	1.2	21.4	19.1 (13.7-25.0)
Women	236	1822	26	14.3 (9.7-20.9)	1.7	0.9	13.4	15.6 (10.2-21.7)
40-49 y	1343	9861	202	20.5 (17.9-23.5)	29.9	3.0	17.5	6.8 (5.9-7.7)
Men	725	5250	117	22.3 (18.6-26.7)	17.6	3.3	18.9	6.7 (5.5-7.9)
Women	618	4611	85	18.4 (14.9-22.8)	12.4	2.7	15.8	6.9 (5.4-8.4)
Stroke Not O	therwise Spec	ified						
Total	3580	27 956	361	12.9 (11.7-14.3)	68.9	2.5	10.5	5.2 (4.7-5.8)
Men	1626	12 551	200	15.9 (13.9-18.3)	37.1	3.0	13.0	5.4 (4.7-6.2)
Women	1954	15 405	161	10.5 (9.0-12.2)	31.8	2.1	8.4	5.1 (4.3-5.9)
18-29 y	213	ND ^d	ND ^d	ND ^d	ND ^d	ND ^d	ND ^d	ND ^d
Men	69	ND ^d	ND ^d	ND ^d	ND ^d	ND ^d	ND ^d	ND ^d
Women	144	ND ^d	ND ^d	ND ^d	ND ^d	ND ^d	ND ^d	ND ^d
30-39 y	769	6145	56	9.1 (7.0-11.8)	6.6	1.1	8.0	8.5 (6.4-10.7)
Men	310	2406	30	12.5 (8.7-17.8)	3.1	1.3	11.2	9.8 (6.6-13.5)
Women	150	3730	26	70(47-102)	3.6	1.0	6.0	7 3 (4 8-10 1)

(continued)

2118 JAMA June 4, 2019 Volume 321, Number 21

Table 2. Long-term Mortality in 30-Day Stroke Survivors Compared With Mortality of the General Population (continued)										
	Total	Patient-Years at Risk	Observed Deaths	Observed Deaths per 1000 Person-Years (95% CI)	Expected Deaths ^a	Expected Deaths per 1000 Person-Years	Excess Rate per 1000 Person-Years ^b	Standardized Mortality Rate (95% CI) ^c		
40-49 y	2598	20 072	296	14.8 (13.2-16.5)	61.5	3.1	11.7	4.8 (4.3-5.4)		
Men	1247	9624	163	16.9 (14.6-19.7	33.7	3.5	13.4	4.8 (4.1-5.6)		
Women	1351	10449	133	12.7 (10.8-15.1)	27.8	2.7	10.1	4.8 (4.0-5.6)		

Abbreviation: ND, not disclosed

^a Expected deaths retrieved from mortality data of the Dutch population matched for age, sex, and calendar year characteristics (Human Mortality Database).

^b The excess mortality rate was calculated as (observed deaths – expected deaths) / person-years at risk expressed per 1000 person-years.

^c Standardized mortality rate is the ratio of the observed mortality rate

divided by expected mortality rate, assuming the observed deaths follow a Poisson distribution. For subgroup analyses within any stroke and the different stroke subtypes, the significance threshold was set to a Bonferroni-adjusted *P* value of .005. *P* < .001 for all rows that have data.

^d Subgroup was too small in number of patients to adequately protect privacy according to legislation regarding the use of this register-based data set of Statistics Netherlands.

hemorrhage (yearly APC, -3.8% [95% CI, -11.9% to 4.4%]; *P* < .001; *R*² = 0.8), and from 4.6% (n = 12) in 1998 to 1.2% (n = 3) in 2010 for stroke not otherwise specified (yearly APC, -0.4% [95% CI, -36.3% to 35.5%; *P* = .006; *R*² = 0.5).

One-year mortality in 30-day survivors after any stroke decreased from 3.7% (n = 39) in 1998 to 2.3% (n = 29) in 2010 (yearly APC, -2.2% [95% CI, -13.0% to 8.6%]; P = .002; $R^2 = 0.6$). In ischemic stroke, 1-year mortality decreased from 3.3% (n = 18) in 1998 to 1.2% (n = 9) in 2010 (yearly APC, -2.8% [95% CI, -20.2% to 14.7%]; P = .005; $R^2 = 0.5$), whereas in intracerebral hemorrhage, it was 7.8% (n = 18) in 1998 and 6.1% (n = 13) in 2010 (yearly APC, 5.4% [95% CI, -18.6% to 29.4%]; P = .07; $R^2 = 0.3$) and in stroke not otherwise specified, 1-year mortality remained stable (3.6% [n = 9] in 1998 and 2.8% [n = 7] in 2010; yearly APC, 24.1% [95% CI, -18.8% to 67.0%]; P = .97; $R^2 = 0.0$).

Cumulative 5-year mortality in 30-day stroke survivors decreased significantly over time in any stroke from 8.3% (n = 89) in 1998 to 5.2% (n = 64) in 2010 (yearly APC, -3.1% [95% CI, -10.2% to 4.0%]; *P* for interaction < .001; R^2 = 0.8), in ischemic stroke from 8.0% (n = 45) to 3.0% (n = 23) (yearly APC, -6.2% [95% CI, -17.2% to 4.8%; *P* for interaction < .001; R^2 = 0.8), and in stroke not otherwise specified from 8.0% (n = 22) to 7.2% (n = 18) (yearly APC, 2.8% to [95% CI, -11.7% to 23.4%]; *P* for interaction = .046; R^2 = 0.3). Cumulative 5-year mortality remained stable in intracerebral hemorrhage (9.6% [n = 23] in 1998 and 12.1% [n = 25] in 2010; yearly APC, 5.9% [95% CI, -13.7% to 19.3%]; *P* for interaction = .98; R^2 = 0.0) (eTable 6 in the Supplement).

From 1998 to 2010, length of stay after any stroke decreased significantly from a mean of 18.2 days to a mean of 8.6 days, resulting in a yearly decrease of -0.8 days (95% CI, -0.2 to 1.3 days; P < .001; $R^2 = 0.4$). This was also true for ischemic stroke, intracerebral hemorrhage, and stroke not otherwise specified.

Causes of Death

Of the 30-day survivors, a total of 1764 patients died during follow-up (13.1%), of which 267 (15.1%) were due to recurrent stroke or were stroke related and 302 (17.1%) due to other cardiovascular diseases. The main cause of death was malignancy (n = 577, 32.7%). The remaining patients (n = 617, 34.9%) died as a result of infection, trauma, and miscellaneous causes. The proportion of deaths attributable to malignancies was higher in the intracerebral hemorrhage group than the group with an ischemic stroke as the index event (41.5% [n = 145] vs 28.8% [n = 273]; $\chi^2 = 18.9$; P < .001). The proportion of cardiovascular-related deaths was higher in the ischemic stroke group than the intracerebral hemorrhage group (19.4% [n = 184] vs 7.2% [n = 25]; $\chi^2 = 28.4$; P < .001). Also, the Poisson regression model showed significant interaction between stroke subtype and malignancies, as well as cardiovascular-related deaths (P < .001), even after correcting for sex in this model (P < .001). The different categories of causes of death stratified by stroke subtype are listed in **Table 4**.

Discussion

Among young adults aged 18 to 49 years in the Netherlands with first stroke, mortality risk compared with the general population remained elevated up to 15 years later.

Major strengths of this study include the populationbased setting with a large number of young patients with stroke (15 257 patients <50 years), whereas in other large studies, patients were included up to 55 years (with substantially fewer patients <50 years). In addition, this populationbased setting increased the likelihood of complete ascertainment of (cause of) death, whereas referral bias has occurred in previous other studies with a hospital-based setting because the more severely affected patients are more likely to die already at home and will not be included in hospitalbased populations.

This study reports, to our knowledge, for the first time the risk and causes of death, cumulative mortality, and annual mortality after intracerebral hemorrhage at young age in large numbers that allow for sufficient power to perform age- and sex-stratified analysis. Another strength of this study is that *ICD* codes for all stroke subtypes were validated in young age groups specifically.

Furthermore, due to the availability of longitudinal data, in combination with the very large numbers, this study was able to report on differential time trends of mortality including case fatality after ischemic stroke and intracerebral hemorrhage.

	Analysis, Hazard	Ratio (95% CI)
	Bivariable ^a	Multivariable
Any Stroke		
Age at index event, y ^c		
25-29	1.5 (0.9-2.6)	1.5 (0.9-2.7)
30-34	1.5 (0.9-2.5)	1.4 (0.9-2.4)
35-39	2.4 (1.5-3.8)	2.3 (1.4-3.6)
40-44	2.7 (1.7-4.2)	2.5 (1.6-4.0)
45-49	3.8 (2.4-5.9)	3.4 (2.1-5.3)
Men ^d	1.5 (1.4-1.7)	1.4 (1.3-1.6)
Charlson Comorbidity Index score ^e		
1	2.1 (1.9-2.4)	2.1 (1.8-2.4)
2	6.5 (5.6-7.5)	5.9 (5.1-6.9)
≥ 3	9.4 (7.5-11.7)	7.7 (6.2-9.7)
Length of stay, d ^f		
3-7	1.1 (0.9-1.3)	1.2 (1.0-1.4)
8-14	1.2 (1.0-1.4)	1.3 (1.1-1.5)
≥15	2.2 (1.9-2.6)	2.2 (1.9-2.6)
schemic Stroke	. ,	
Age at index event, y ^c		
25-29	1.5 (0.7-3.1)	1.4 (0.6-2.9)
30-34	0.9 (0.4-1.8)	0.8 (0.4-1.7)
35-39	1.9 (1.0-3.7)	1.9 (1.0-3.7)
40-44	2.4 (1.3-4 5)	2.2 (1.2-4 2)
45-49	3.5 (1.9-6.5)	3.1 (1.7-5.8)
/len ^d	1.5 (1.4-1.8)	1.4 (1.3-1.6)
harlson Comorbidity ndex score ^e		(, ,)
1	2.1 (1.7-2.5)	2.0 (1.7-2.4)
2	6.3 (5.1-7.7)	5.5 (4.5-6.8)
≥ 3	9.7 (7.1-13.4)	7.6 (5.5-10.4
ength of stay, d ^f		
3-7	1.2 (0.9-1.5)	1.3 (1.0-1.7)
8-14	1.2 (1.0-1.6)	1.3 (1.0-1.7)
≥15	2.3 (1.8-2.9)	2.4 (1.9-3.0)
ntracerebral Hemorrhage		
Age at index event, y ^c		
25-29	2.7 (0.9-7.5)	3.2 (1.1-9.1)
30-34	3.0 (1.2-7.9)	2.8 (1.1-7.2)
35-39	4.5 (1.8-11.2)	3.6 (1.4-9.0)
40-44	4.4 (1.8-10.7)	4.0 (1.6-9.9)
45-49	5.1 (2.1-12.4)	4.6 (1.9-11.3
Men ^d	1.4 (1.1-1.7)	1.4 (1.1-1.7)
Charlson Comorbidity ndex score ^e	. ,	. ,
1	1.7 (1.3-2.2)	1.7 (1.3-2.3)
2	8.2 (6.1-11.2)	8.3 (6.1-11.2
≥3	9.5 (6.1-15.0)	9.3 (5.9-14.8
Length of stay, d ^f		
3-7	1.2 (0.8-1.8)	1.3 (0.8-2.0)
	. ,	

(continued)

1.3 (0.8-1.9)

2.1 (1.4-3.0)

Table 3. Factors Associated With Mortality Among 30-Day Stroke Survivors According to Stroke Subtype (continued)

	Analysis, Hazard R	atio (95% CI)		
	Bivariable ^a	Multivariable ^b		
Stroke Not Otherwise Specified				
Age at index event, y ^c				
25-29	0.7 (0.2-2.5)	0.7 (0.2-2.7)		
30-34	1.7 (0.6-5.0)	1.7 (0.6-4.8)		
35-39	2.0 (0.7-5.5)	1.7 (0.6-4.8)		
40-44	2.2 (0.8-6.0)	2.0 (0.7-5.3)		
45-49	3.4 (1.3-9.0)	2.6 (1.0-7.1)		
Men ^d	1.5 (1.2-1.8)	1.4 (1.2-1.7)		
Charlson Comorbidity Index score ^e				
1	2.5 (2.0-3.1)	2.4 (1.9-3.0)		
2	5.5 (4.1-7.4)	5.0 (3.7-6.8)		
≥3	7.7 (5.0-11.9)	6.3 (4.1-9.8)		
Length of stay,d ^f				
3-7	1.0 (0.7-1.3)	1.1 (0.8-1.4)		
8-14	1.1 (0.8-1.5)	1.2 (0.9-1.4)		
≥15	2.1 (1.6-2.7)	2.1 (1.6-2.8)		

^a Hazard ratios (95% CI) were computed separately for age at index event, sex, Charlson Comorbidity Index score, and length of stay.

^b Age at index event, sex, Charlson Comorbidity Index score, and length of stay were entered simultaneously in a Cox proportional hazards model.

^c Reference category for age at index event is 18-24 years.

^d Reference category for sex is women.

^e Reference category for Charlson Comorbidity Index is a score of 0. Charlson Comorbidity Index score ranges from 0 to 6, with higher scores indicating more comorbidity.

^f Reference category for length of stay is 0-2 days.

Additionally, this study reports on outcomes of patients who were treated with stroke unit care, intravenous thrombolysis, and hemicraniectomy and secondary preventive treatment, whereas earlier studies included patients who had their stroke years before the implementation of these therapies.²³⁻²⁵

In addition, within this large cohort, it was possible to analyze specific causes of death in more detail than previous studies, with information available on subtypes of malignancies. Previous studies have presented this only for patients with ischemic stroke and not for patients with intracerebral hemorrhage, and causes of death were limited to recurrent stroke, other cardiovascular disease, malignancies, and other causes.

A limited number of earlier studies with a comparable duration of follow-up has shown comparable mortality rates of 4.6% (3 year) and 16% (16 year).^{7,9,10,15} Most previous studies on long-term mortality after stroke in young adults were hospital based, had varying inclusion criteria, or also included SAH or TIA. These disorders may have a different prognosis than ischemic stroke and intracerebral hemorrhage and may, therefore, bias the mortality rates after stroke. In young patients with TIA, the 1-year cumulative mortality was shown to be only 0.7% and in the years thereafter, 0.2% per year.^{7,9,10} In contrast, patients with an SAH had a 17% excess mortality rate compared with the general population after 20 years of follow-up.²⁶

8-14

≥15

1.2(0.8-1.9)

2.0 (1.4-2.9)

	Index Event, No. (%)											
	Any Stroke			Ischemic Stroke			Intracerebral Hemorrhage			Stroke Not Otherwise Specified		
Causes of Death ^a	Total	Men	Women	Total	Men	Women	Total	Men	Women	Total	Men	Women
Total	1763 (100)	987 (100)	776 (100)	947 (100)	524 (100)	423 (100)	349 (100)	210 (100)	139 (100)	467 (100)	253 (100)	214 (100)
Stroke related	267 (15.1)	147 (14.9)	120 (15.5)	129 (13.6)	72 (13.7)	57 (13.5)	79 (22.6)	42 (20.0)	37 (26.6)	59 (12.6)	33 (13.0)	26 (12.1)
Ischemic stroke	35 (2.0)	20 (2.0)	15 (1.9)	28 (3.0)	ND^{b}	ND^{b}	ND^{b}	ND^{b}	ND^{b}	ND^b	ND^{b}	ND^{b}
Intracerebral hemorrhage	49 (2.8)	26 (2.6)	23 (3.0)	12 (1.3)	ND^{b}	ND^{b}	ND^{b}	ND^{b}	ND^{b}	ND^b	ND^b	ND ^b
Other stroke-related deaths	183 (10.4)	101 (10.2)	82 (10.6)	89 (9.4)	50 (9.5)	39 (9.2)	50 (14.3)	26 (12.4)	24 (17.3)	44 (9.4)	25 (9.9)	19 (8.9)
Cardiac and other vascular events	302 (17.1)	194 (19.7)	108 (13.9)	184 (19.4)	113 (21.6)	71 (16.8)	25 (7.2)	ND ^b	ND ^b	93 (19.9)	63 (24.9)	30 (14.0)
Malignancies	577 (32.7)	273 (27.7)	304 (39.2)	273 (28.8)	127 (24.2)	146 (34.5)	145 (41.5)	80 (38.1)	65 (46.8)	159 (34.0)	92 (26.1)	126 (43.5)
Lung cancer	145 (25.1)	53 (19.4)	92 (30.3)	83 (30.4)	31 (24.4)	52 (35.6)	14 (9.7)	ND^{b}	ND^{b}	48 (30.2)	ND ^b	ND ^b
Brain tumor	86 (14.9)	56 (20.5)	30 (9.9)	29 (10.6)	16 (12.6)	13 (8.9)	44 (30.3)	32 (40.0)	12 (18.5)	13 (8.2)	34 (51.5)	38 (40.9)
Hematological malignancies	47 (8.1)	22 (8.1)	25 (8.2)	ND ^b	ND ^b	ND ^b	ND ^b	ND ^b	ND ^b	ND ^b	ND ^b	ND ^b
Breast cancer	40 (6.9)	ND^{b}	ND^b	22 (8.1)	ND^b	ND^b	ND ^b	ND ^b	ND^{b}	12 (7.5)	ND^{b}	ND ^b
Melanoma	28 (4.9)	ND ^b	21 (14.5)	10 (12.5)	11 (16.9)	ND ^b	ND ^b	ND ^b				
Other forms of malignancies	232 (40.2)	128 (46.9)	103 (33.9)	115 (42.1)	71 (55.9)	44 (30.1)	46 (31.7)	24 (30.0)	22 (33.8)	70 (44.0)	33 (50.0)	37 (39.8)
Infections	87 (4.9)	50 (5.1)	37 (4.8)	52 (5.5)	32 (6.1)	20 (4.7)	19 (5.4)	ND ^b	ND^{b}	16 (3.4)	ND ^b	ND ^b
Trauma	78 (4.4)	53 (5.4)	25 (3.2)	42 (4.4)	29 (5.5)	13 (3.1)	12 (3.4)	ND ^b	ND ^b	24 (5.1)	ND ^b	ND ^b
Miscellaneous	452 (25.6)	270 (27.4)	182 (23.5)	267 (28.2)	151 (28.8)	116 (27.4)	69 (19.8)	50 (23.8)	19 (13.7)	116 (24.8)	69 (27.3)	47 (22.0)

Table 4. Causes of Death in 30-Day Survivors, Stratified by Stroke Subtype and Sex

Abbreviation: ND, not disclosed.

^a International Classification of Diseases codes used to define the categories of causes of death are listed in eTable 3 in the Supplement. ^b Subgroup was too small in the number of patients to adequately protect privacy according to legislation regarding the use of this register-based data set of Statistics Netherlands.

The excess long-term mortality in young adults following stroke compared with age- and sex-matched individuals in the general population may suggest that even after treatment for stroke and treatment of associated risk factors according to current standards, the risk of death of young patients with stroke remained increased compared with their population peers. From 1998 to 2017, the already low risk of death in young adults in the general population has continued to decrease, possibly because of improved treatment of other life-threatening diseases, such as malignancies, and because of fewer traffic crashes.²⁷

The observed decreasing trend observed in mortality after ischemic stroke in young adults could be partially due to improved diagnosis of ischemic stroke with better and more imaging techniques available. In this way, minor and resolved syndromes can be increasingly diagnosed as stroke (by magnetic resonance imaging), which would partially explain decreasing rates of poststroke mortality only for ischemic stroke. This hypothesis is supported by the fact that the 1- and 5-year mortality rates of intracerebral hemorrhage did not decrease significantly over the study period, which can be readily diagnosed purely by use of a computed tomographic scan, which was already available in the earlier years of the study period.

A previous study also found male sex as a risk factor for long-term mortality.⁹ The increased risk of death in men may be due to differences in risk factors and etiology between men and women (eg, higher prevalence of traditional vascular risk factors and more large-artery disease in men²⁸⁻³⁰) as the incidence of stroke in the study period was higher in young women.⁴ When compared with the general population, no significant differences between men and women in their risk of death were found.

In Sweden, a decline in case fatality of ischemic stroke was seen in men aged 30 to 84 years, but not in women. A similar difference with significant decrease of case fatality only in men was seen in the Framingham study.^{31,32} The decrease in case fatality of ischemic stroke this study found for both young men and women might be attributable to both the change in the organization of stroke care with better education, introduction of stroke units in the early 1990s, the introduction of hemicraniectomy for space-occupying infarctions,^{23,33} and the higher detection of more minor and resolved syndromes as mentioned here. The reason that case fatality of intracerebral hemorrhage has remained stable over time may be the limited treatment options for this type of stroke compared with ischemic stroke.³³⁻³⁶

This study explored the causes of death in young adults after stroke, which may provide evidence for possible underlying disease mechanisms. For example, a higher percentage of patients who died of malignancies and cardiovascular

disease was found compared with the general population. Malignancies were responsible for 32.7% of deaths in this cohort, whereas in corresponding age groups in published records of Statistics Netherlands, 25.6% died of malignancy in the period from 1998 to 2010.³⁷ A total of 17.1% of patients died of cardiovascular diseases, whereas in corresponding age groups from the general population, this percentage was 10.6%, according to Statistics Netherlands.³⁷ This may suggest that the underlying risk factors and causes of stroke continue to expose patients to new events throughout the rest of their lives.¹³ However, results of causes of death that are based on small patient numbers should be interpreted with caution.

Limitations

This study has several limitations. First, due to the registrybased study design, there was an inability to control for possible confounders (eg, stroke severity, family history, medication) and comorbidity could only be assessed with the CCI, which provides a reliable measure of someone's comorbidity, but without information about actual risk factors and underlying etiology.²⁰ Specifically for young patients with stroke, with a wide variety of risk factors and causes, more detailed information would have been desirable. Because the CCI is composed of comorbidities defined by a previous hospital admission before the time of index event, this may underestimate the effect of comorbidity on long-term mortality when comorbidity increases between index stroke and death. Conversely, the adjustment for baseline CCI score may result in overestimation when comorbidity decreases after the index stroke until death. This possible bias is expected to be very low because young adults are less likely to have major comorbidity than elderly patients.

Second, because fewer hospitals contributed to the HDR register from 2006 onwards, the incidence of stroke could have been underestimated. If these records were not missing completely at random, some bias may have been introduced into the analysis.

Third, strokes between 1998 and 2010 were defined as first if no earlier admission for stroke was registered between 1995 and 2010. A limitation of this method is that a very small proportion of incident strokes may have been misclassified as "first," if these patients would have had a stroke before 1995.³⁸ However, given the 2% yearly risk of recurrent stroke 3 years after index stroke, this risk of misclassification is considered very low.³⁸

Fourth, using the general population as a control group includes young adults with stroke, which could possibly bias the study toward the null. However, because only a very small proportion of the controls will have had a stroke (in the order of a few cases per thousands of controls), this is unlikely to have had a major effect on the findings.

Fifth, given that recent advances in the management of ischemic stroke with mechanical thrombectomy occurred largely after the end date of the data included in this study, the risk of mortality after ischemic stroke may not entirely be generalizable to reflect contemporary management.

Conclusions

Among young adults aged 18 to 49 years in the Netherlands who were 30-day survivors of first stroke, mortality risk compared with the general population remained elevated up to 15 years later.

ARTICLE INFORMATION

Accepted for Publication: April 30, 2019. Published Online: May 23, 2019.

doi:10.1001/jama.2019.6560

Author Contributions: Drs de Leeuw and Klijn had full access to all of the data in the study and take responsibility for the integrity of the data and the accuracy of the data analysis. Drs Ekker and Verhoeven contributed equally. Drs Klijn and de Leeuw contributed equally. *Concept and design:* Ekker, Verhoeven, Klijn,

de Leeuw.

Acquisition, analysis, or interpretation of data: All authors.

Drafting of the manuscript: Ekker, Verhoeven. Critical revision of the manuscript for important intellectual content: All authors. Statistical analysis: Ekker, Verhoeven. Obtained funding: Klijn, de Leeuw. Administrative, technical, or material support: Ekker, Verhoeven, Klijn. Supervision: Klijn, de Leeuw.

Conflict of Interest Disclosures: Dr Klijn reported receiving grants from Netherlands Cardiovascular Research Initiative, which is supported by the Dutch Heart Foundation, CVON2015-01: CONTRAST, and the Brain Foundation Netherlands (HA2015-01-06). She also reported receiving grants from the Dutch Heart Foundation (2012T077) and the Netherlands Organization for Health Research and Development (ZonMw; grant 015008048) outside the submitted work. No other disclosures were reported.

Funding/Support: Dr Vaartjes is funded by the Dutch Heart Foundation for project Facts and Figures. Dr Klijn is supported by a clinical established investigator grant from the Dutch Heart Foundation (2012T077) and by an ASPASIA grant from the Netherlands Organisation for Health Research and Development (ZonMw, grant 015008048). Dr de Leeuw is supported by a clinical established investigator grant from the Dutch Heart Foundation (2014 T060) and by a VIDI innovational grant from the Netherlands ZonMw (grant 016126351).

Role of the Funder/Sponsor: The funders had no role in the design and conduct of the study; collection, management, analysis, and interpretation of the data; preparation, review, or approval of the manuscript; and decision to submit the manuscript for publication.

REFERENCES

 Barker-Collo S, Bennett DA, Krishnamurthi RV, et al; GBD 2013 Writing Group; GBD 2013 Stroke Panel Experts Group. Sex differences in stroke incidence, prevalence, mortality and disability-adjusted life years: results from the Global Burden of Disease Study 2013. *Neuroepidemiology*. 2015;45(3):203-214. doi:10.1159/000441103

2. Feigin VL, Norrving B, Mensah GA. Global burden of stroke. *Circ Res.* 2017;120(3):439-448. doi:10.1161/CIRCRESAHA.116.308413

3. World Health Organization. Cardiovascular diseases (CVDs). https://www.who.int/news-room/ fact-sheets/detail/cardiovascular-diseases-(cvds). Accessed May 10, 2019.

4. Ekker MS, Verhoeven JI, Vaartjes I, van Nieuwenhuizen KM, Klijn CJM, de Leeuw FE. Stroke-incidence in young adults according to age, subtype, sex, and time-trends: a nationwide registry-based study. *Neurology*. 2019;92:1-11. doi: 10.1212/WNL.000000000007533

 Béjot Y, Daubail B, Jacquin A, et al. Trends in the incidence of ischaemic stroke in young adults between 1985 and 2011: the Dijon Stroke Registry. *J Neurol Neurosurg Psychiatry*. 2014;85(5):509-513. doi:10.1136/jnnp-2013-306203

6. Sultan S, Elkind MS. The growing problem of stroke among young adults. *Curr Cardiol Rep.* 2013; 15(12):421. doi:10.1007/s11886-013-0421-z

7. Waje-Andreassen U, Naess H, Thomassen L, Eide GE, Vedeler CA. Long-term mortality among young ischemic stroke patients in western Norway. *Acta Neurol Scand*. 2007;116(3):150-156. doi:10.1111/j. 1600-0404.2007.00822.x

8. Putaala J, Curtze S, Hiltunen S, Tolppanen H, Kaste M, Tatlisumak T. Causes of death and predictors of 5-year mortality in young adults after first-ever ischemic stroke: the Helsinki Young Stroke Registry. *Stroke*. 2009;40(8):2698-2703. doi:10.1161/STROKEAHA.109.554998

9. Naess H, Waje-Andreassen U. Review of long-term mortality and vascular morbidity amongst young adults with cerebral infarction. *Eur J Neurol*. 2010;17(1):17-22. doi:10.1111/j.1468-1331. 2009.02868.x

10. Koivunen RJ, Tatlisumak T, Satopää J, Niemelä M, Putaala J. Intracerebral hemorrhage at young age: long-term prognosis. *Eur J Neurol*. 2015;22(7): 1029-1037. doi:10.1111/ene.12704

11. Koton S, Schneider AL, Rosamond WD, et al. Stroke incidence and mortality trends in US communities, 1987 to 2011. *JAMA*. 2014;312(3):259-268. doi:10.1001/jama.2014.7692

12. Marini C, Totaro R, De Santis F, Ciancarelli I, Baldassarre M, Carolei A. Stroke in young adults in the community-based L'Aquila registry: incidence and prognosis. *Stroke*. 2001;32(1):52-56. doi:10. 1161/01.STR.32.1.52

13. Rutten-Jacobs LC, Arntz RM, Maaijwee NA, et al. Long-term mortality after stroke among adults aged 18 to 50 years. *JAMA*. 2013;309(11):1136-1144. doi:10.1001/jama.2013.842

14. Nieuwkamp DJ, Vaartjes I, Algra A, Bots ML, Rinkel GJ. Age- and gender-specific time trend in risk of death of patients admitted with aneurysmal subarachnoid hemorrhage in the Netherlands. *Int J Stroke*. 2013;8(suppl A100):90-94. doi:10.1111/ijs. 12006

15. Giang KW, Björck L, Nielsen S, et al. Twenty-year trends in long-term mortality risk in 17,149 survivors of ischemic stroke less than 55 years of age. *Stroke*. 2013;44(12):3338-3343. doi:10.1161/STROKEAHA.113. 002936

16. Vaartjes I, Reitsma JB, de Bruin A, et al. Nationwide incidence of first stroke and TIA in the Netherlands. *Eur J Neurol*. 2008;15(12):1315-1323. doi:10.1111/j.1468-1331.2008.02309.x

17. Vaartjes I, O'Flaherty M, Capewell S, Kappelle J, Bots M. Remarkable decline in ischemic stroke mortality is not matched by changes in incidence. *Stroke*. 2013;44(3):591-597. doi:10.1161/STROKEAHA. 112.677724

 Jolink WMT, Klijn CJ, Brouwers PJ, Kappelle LJ, Vaartjes I. Time trends in incidence, case fatality, and mortality of intracerebral hemorrhage. *Neurology*. 2015;85(15):1318-1324. doi:10.1212/WNL. 00000000002015 **19**. Tirschwell DL, Longstreth WT Jr. Validating administrative data in stroke research. *Stroke*. 2002;33(10):2465-2470. doi:10.1161/01.STR. 0000032240.28636.BD

20. Quan H, Li B, Couris CM, et al. Updating and validating the Charlson Comorbidity Index and score for risk adjustment in hospital discharge abstracts using data from 6 countries. *Am J Epidemiol.* 2011;173(6):676-682. doi:10.1093/aje/kwq433

21. Barbieri M, Wilmoth JR, Shkolnikov VM, et al. Data resource profile: the Human Mortality Database (HMD). *Int J Epidemiol*. 2015;44(5):1549-1556. doi:10.1093/ije/dyv105

22. Harteloh P, de Bruin K, Kardaun J. The reliability of cause-of-death coding in the Netherlands. *Eur J Epidemiol*. 2010;25(8):531-538. doi:10.1007/ s10654-010-9445-5

23. Hofmeijer J, Kappelle LJ, Algra A, Amelink GJ, van Gijn J, van der Worp HB; HAMLET investigators. Surgical decompression for space-occupying cerebral infarction (the Hemicraniectomy After Middle Cerebral Artery infarction with Life-threatening Edema Trial [HAMLET]): a multicentre, open, randomised trial. *Lancet Neurol*. 2009;8(4):326-333. doi:10.1016/S1474-4422(09) 70047-X

24. Shepherd J, Blauw GJ, Murphy MB, et al; PROSPER study group. PROspective Study of Pravastatin in the Elderly at Risk. Pravastatin in elderly individuals at risk of vascular disease (PROSPER): a randomised controlled trial. *Lancet*. 2002;360(9346):1623-1630. doi:10.1016/ S0140-6736(02)11600-X

25. Mihaylova B, Emberson J, Blackwell L, et al; Cholesterol Treatment Trialists' (CTT) Collaborators. The effects of lowering LDL cholesterol with statin therapy in people at low risk of vascular disease: meta-analysis of individual data from 27 randomised trials. *Lancet*. 2012;380(9841):581-590. doi:10.1016/S0140-6736(12)60367-5

26. Huhtakangas J, Lehto H, Seppä K, et al. Long-term excess mortality after aneurysmal subarachnoid hemorrhage: patients with multiple aneurysms at risk. *Stroke*. 2015;46(7):1813-1818. doi:10.1161/STROKEAHA.115.009288

27. National Agency for Waterways and Public Works. Number of road deaths stable in 2014. https://www.cbs.nl/en-gb/news/2015/18/numberof-road-deaths-stable-in-2014. Published July 5, 2015. Accessed November 14, 2018.

28. Béjot Y, Bailly H, Durier J, Giroud M. Epidemiology of stroke in Europe and trends for the 21st century. *Presse Med*. 2016;45(12 pt 2):e391e398. doi:10.1016/j.lpm.2016.10.003 **29**. Putaala J, Metso AJ, Metso TM, et al. Analysis of 1008 consecutive patients aged 15 to 49 with first-ever ischemic stroke: the Helsinki Young Stroke Registry. *Stroke*. 2009;40(4):1195-1203. doi: 10.1161/STROKEAHA.108.529883

30. Tibæk M, Dehlendorff C, Jørgensen HS, Forchhammer HB, Johnsen SP, Kammersgaard LP. Increasing incidence of hospitalization for stroke and transient ischemic attack in young adults: a registry-based study. *J Am Heart Assoc.* 2016;5(5): e003158. doi:10.1161/JAHA.115.003158

31. Vangen-Lønne AM, Wilsgaard T, Johnsen SH, Carlsson M, Mathiesen EB. Time trends in incidence and case fatality of ischemic stroke: the tromsø study 1977-2010. *Stroke*. 2015;46(5):1173-1179. doi: 10.1161/STROKEAHA.114.008387

32. Carandang R, Seshadri S, Beiser A, et al. Trends in incidence, lifetime risk, severity, and 30-day mortality of stroke over the past 50 years. *JAMA*. 2006;296(24):2939-2946. doi:10.1001/jama.296. 24.2939

 Schreuder FH, Sato S, Klijn CJ, Anderson CS. Medical management of intracerebral haemorrhage. *J Neurol Neurosurg Psychiatry*. 2017; 88(1):76-84. doi:10.1136/jnnp-2016-314386

34. Brainin M, Olsen TS, Chamorro A, et al; EUSI Executive Committee; EUSI Writing Committee; European Stroke Initiative. Organization of stroke care: education, referral, emergency management and imaging, stroke units and rehabilitation. *Cerebrovasc Dis.* 2004;17(suppl 2):1-14. doi:10.1159/000074816

35. Asadi H, Dowling R, Yan B, Wong S, Mitchell P. Advances in endovascular treatment of acute ischaemic stroke. *Intern Med J.* 2015;45(8):798-805. doi:10.1111/imj.12652

36. Aarnio K, Haapaniemi E, Melkas S, Kaste M, Tatlisumak T, Putaala J. Long-term mortality after first-ever and recurrent stroke in young adults. *Stroke*. 2014;45(9):2670-2676. doi:10.1161/STROKEAHA.114. 005648

37. Population; generation, gender, age and origin. Statistics Netherlands. CBS Statline. http://statline.cbs.nl/statweb/publication/?DM= SLNL&PA=37325&D1=0&D2=a&D3=0&D4=a&D6= 18-21&VW=T. Updated July 18, 2017. Accessed January 26, 2017.

38. Rutten-Jacobs LC, Maaijwee NA, Arntz RM, et al. Long-term risk of recurrent vascular events after young stroke: the FUTURE study. *Ann Neurol.* 2013;74(4):592-601. doi:10.1002/ana.23953